# Cancer Risk Among Firefighters: A Review and Meta-analysis of 32 Studies 

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Objective: The objective of this study was to review 32 studies on firefighters and to quantitatively and qualitatively determine the cancer risk using a meta-analysis. Methods: A comprehensive search of computerized databases and bibliographies from identified articles was performed. Three criteria used to assess the probable, possible, or unlikely risk for 21 cancers included pattern of meta-relative risks, study type, and heterogeneity testing. Results: The findings indicated that firefighters had a probable cancer risk for multiple myeloma with a summary risk estimate (SRE) of 1.53 and $95 \%$ confidence interval (CI) of 1.21-1.94, non-Hodgkin lymphoma $(S R E=1.51,95 \% C I=1.31-1.73)$, and prostate $(S R E=1.28 ; 95 \% C I=1.15-1.43)$. Testicular cancer was upgraded to probable because it had the highest summary risk estimate (SRE $=2.02 ; 95 \%$ $C I=1.30-3.13$ ). Eight additional cancers were listed as having a "possible" association with firefighting. Conclusions: Our results confirm previous findings of an elevated metarelative risk for multiple myeloma among firefighters. In addition, a probable association with non-Hodgkin lymphoma, prostate, and testicular cancer was demonstrated. (J Occup Environ Med. 2006;48: 1189-1202)

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uring the course of their work, firefighters are exposed to harmful substances at the fire scene as well as at the firehouse. At the fire scene, firefighters are potentially exposed to various mixtures of particulates, gases, mists, fumes of an organic and/or inorganic nature, and the resultant pyrolysis products. ${ }^{1,2}$ Specific potential exposures include metals such as lead, antimony, cadmium, uranium, chemical substances, including acrolein, benzene, methylene chloride, polyaromatic hydrocarbons, perchlorethylene, toluene, trichloroethylene, trichlorophenol, xylene, formaldehydes, minerals such as asbestos, crystalline, and noncrystalline silica, silicates, and various gases that may have acute, toxic effects. ${ }^{1,2}$ In some situations, respiratory protection equipment may be inadequate or not felt to be needed resulting in unrecognized exposure. ${ }^{3}$ At the firehouse where firefighters spend long hours, exposures may occur to complex mixtures that comprise diesel exhaust, particularly if trucks are run in closed houses without adequate outside venting. In light of the World Trade Center disaster, concerns have reemerged and heightened related to building debris particle exposures from pulverized cement and glass, fiberglass, asbestos, silica, heavy metals, soot, and/or organic products of combustion. ${ }^{3}$

To date, only one meta-analysis conducted by Howe and Burch in 1990 examined the extent of cancer risk among firefighters in 11 mortality studies. ${ }^{4}$ They reported that there was an increased association with the occurrence of brain tumors, malignant melanoma, and multiple myeloma with the evidence in favor of
causality somewhat greater for brain tumors and multiple myeloma. Since then, there have been numerous mortality and incidence studies. Hence, the purpose of this study was twofold. The first purpose was to update the Howe and Burch findings by reviewing the methodologic characteristics of these studies and determining the probability of cancer by assessing the weight of evidence, including the calculated metarisk estimates. The second purpose was to describe a methodology for use in a meta-analysis when diverse investigations are being evaluated and summarized.

## Materials and Methods

## Search Strategy and Inclusion Criteria

Standardized mortality ratio (SMR), proportional mortality ratio (PMR), relative risk (RR), standardized incidence ratio (SIR), and case-control/ mortality odds ratio (OR) studies related to firefighters and cancer risk were evaluated. For publication selection, at least 1 year in service as firefighters was required except for those studies basing employment on death certificates. Publications were retrieved by a search of computerized databases, including Medline (1966-December 2003), Health and Safety Science Abstracts (since 1980-December 2003), Cancerlit (1963-December 2003), NIOSHTIC and NIOSHTIC2 (up to December 2003), BIOSIS Previews (1980December 2003), and PubMed (up to December 2003) using the following key words: firefighters, fire fighters, cancer. In addition to the computerized search, bibliographies in identified papers were reviewed for additional studies.

The search was restricted to reports published in English; abstracts and reviews were not included. Studies were excluded without basic data (eg, confidence intervals) that are necessary in the derivation of the meta-analysis risk estimate. If there was more than one article with the same or overlapping population, preference was given to the article providing more comprehensive information. The
data were extracted from each article by one reviewer and was verified by another. Discrepancies identified by the second reviewer were resolved in a consensus meeting.

Likelihood of Cancer Risk. Statistically significant increases in cancer risks among firefighters were evaluated as the likelihood for cancer risk given a three-criteria assessment. The three criteria included "pattern of meta-relative risk association," "study type," and "consistency" among studies. These criteria were particularly important given the different methodologies used for evaluating cancer risk
(ie, SMR, PMR, RR, SIR, and OR). These criteria were used in a forward approach as illustrated in Figure 1 in which at each stage, a new criterion was applied, and the probability of cancer risk was reassessed. The likelihood for cancer risk was given an assignment of "probable," "possible," or "not likely" patterned after the International Agency for Research on Cancer (IARC) risk assessment of human carcinogenicity in terms of weight of the evidence. ${ }^{5}$

The "pattern of metarelative risk associations" was the first criterion and included a two-step evaluation. For the


Fig. 1. Likelihood of cancer risk.
first step, the strength of the metaanalysis by each study type (eg, SMR, PMR) was assigned a score. The score of " ++ " was assigned if the metarelative risk was statistically significant and greater than 1.1. The score of " + " was assigned if the metarelative risk was not statistically significant, but the point risk estimate was greater than 1.1. The score of "-" was assigned if the metarelative risk was not statistically significant, and the point risk estimate was equal to or less than 1.1. At the second step, these scores were used to assign a probable, possible, or unlikely designation for the pattern of metarelative risk association. A "probable" was assigned to the cancerspecific site if one metarelative risk (ie, mSMR, mPMR, mSMR and PMR, mRR, mSIR, mOR) was statistically significant (score of ++ ) and at least another was greater than 1.1 (score of + ). A "possible" assignment was given if only one metarelative risk was available and was statistically significant (score of ++ ) or if at least two metarelative risks were greater than 1.1 but were not statistically significant (score of + ). "Not likely" was assigned if the cancer-specific site did not meet the probable or possible criteria.

The second criterion examined the "study type" used to generate metarelative risks. If the metarelative risk estimate reached statistical significance (score of ++ ), based primarily on PMR studies, the level was downgraded. PMR studies do not measure the risk of death or death rates but rather the relative frequency of that particular cause among all causes of death. Hence, the limitation of a PMR study is that the estimate may be abnormally low or high based on the overall increase or decrease in mortality and not due to the cause of interest. ${ }^{6}$ Also, if the mSMR point risk estimate was not significant and $\leq 1.1(-)$, the level was downgraded. The third criterion used for generating the likelihood of cancer risk was an assessment of "inconsistency" among studies. Heterogeneity testing as described in statistical methods was used to evaluate
inconsistency. The level was downgraded if heterogeneity (inconsistency) testing among all combined studies had an $\alpha \leq 0.10$.

## Statistical Methods

For all cancer outcomes having two or more studies, the observed and expected values from each study were summed and a metarelative risk estimate (mRR) was calculated. An mRR was calculated for each cancer by each study type, eg, SMR studies and as a summary metarelative risk across all study types. The mRR was defined as the ratio of the total number of observed deaths or incident cases to the total number of expected deaths or incident cases as follows:

$$
m R R=\frac{\sum_{i=1}^{n} O_{i}}{\sum_{i=1}^{n} E_{i}}
$$

where $O_{i}$ denotes observed deaths (cases) in each individual study, $E_{i}$ denotes expected deaths (cases), and $n$ is the total number of studies. ${ }^{7}$ The $95 \%$ confidence interval (CI) of mRR may be computed using the Poisson probability distribution as described by Breslow and Day. ${ }^{8}$ The standard error (SE) for the metarelative risk is calculated as $S E=\frac{1}{\sqrt{\sum W_{i}}}$ where $W_{i}$ is the statistical weight for a given study defined as $1 / S E_{i}^{2}$ and $S E_{i}$ is the standard error for a given study.

In the absence of heterogeneity, the fixed-effect model was applied for deriving the metarelative risk estimate; otherwise, the random-effects model was used. A test for heterogeneity for the fixed-effect approach is given by $Q=\sum_{i=1}^{n} W_{i} *\left\{\log \left(R R_{i}\right)-\log (m R R)\right\}^{2}$ where $R R_{i}$ and $m R R$ are the relative risk and the metarelative risk, respectively. The hypothesis of homogeneity among studies would be rejected if $Q$ exceeds $\chi_{n-1, \alpha}^{2}$. Then the randomeffects model was used with a different study weight $\left(W_{i}^{*}\right)$ that further accounts for the interstudy variation in
effect size. ${ }^{8}$ The weighing factor $W_{i}^{*}$ in the DerSimonian and Laird randomeffects model is

$$
W_{i}^{*}=\frac{1}{\left[D+\left(\frac{1}{w_{i}}\right)\right]}
$$

where $W_{i}$ is the statistical weight for a given study for the fixed-effect model and is equal to $1 / S E_{i}^{2}$ with $S E_{i}$ being the standard error for a given study according to Chen and Seaton ${ }^{9}$

$$
D=\frac{[Q-(n-1)] * \sum_{i=1}^{n} W_{i}}{\left(\sum_{i=1}^{n} W_{i}\right)^{2}-\sum_{i=1}^{n} W_{i}^{2}}
$$

It should be noted that $D$ is set to 0 if $Q<n-1$. The random-effects model was validated against data provided in Petitti, ${ }^{10}$ which after application using our equations gave identical results. For this study, an $\alpha \leq 10 \%$ or less for declaring heterogeneity was adopted. ${ }^{11}$

The SAS software was used to perform the calculations and validated our program for the fixed-effect model using data from different studies compiled by Howe and Burch ${ }^{4}$ on standardized mortality ratios and proportional mortality ratios among firefighters. Where there were no observed deaths or incident cases, the lower confidence interval for an individual study was set at 0.1 as suggested in the method used by Collins and Acquavella. ${ }^{12}$ This method was compared with the data excluding studies with a zero relative risk, and the results were similar.

## Results

## Identification and Characteristics of Studies

The computerized literature search identified 21 U.S. and 14 non-U.S. articles. ${ }^{13-47}$ It was determined that three studies were not eligible for the meta-analysis because of either insufficient data, ${ }^{41}$ data were combined for firefighters and other personnel, ${ }^{42}$ or
the text was not published in English. ${ }^{43}$ In addition, four studies ${ }^{44-47}$ were excluded because of overlapping populations with other reports. ${ }^{18,30}$ For example, in 1992, Demers et al ${ }^{18}$ reported more observed and expected cancers than in the 1994 article. ${ }^{46}$ Four additional studies ${ }^{48-51}$ were identified in the review by Howe and Burch ${ }^{4}$ and used in the meta-analysis. These latter four studies are not presented in Table 1. Hence, a total of 28 studies received a detailed review as shown in Table 1, which describes the study design characteristics, exposure, and outcome definitions. Sixteen were U.S. studies and 12 were non-U.S. investigations. Five studies had an internal comparison group with the remaining using regional or national comparison groups. Fourteen ascertained exposures from employment records and defined exposure as a dichotomous (yes/no) variable. The majority of the studies relied on death certificates for assessing a cancer diagnosis. Of a total of 32 articles, 26 are included in the metaadditional articles are case-control/ mortality odds ratio studies and presented in Table 3 with one meta- analysis for non-Hodgkin's lymphoma.

## Overview of Meta-analysis

Table 2 summarizes the metaanalysis results by study type. Studies were mostly mortality and were analyzed using SMRs and PMRs. All-cause mortality had an SMR $10 \%$ less than general population rates. Mortality from all cancers was similar to the general population using SMR and RR indices, but PMR studies showed a $10 \%$ significantly higher rate (Table 2). For individual cancers, there were statistically significant elevated meta-SMR estimates for colon cancer (1.34) and multiple myeloma (1.69). PMR studies demonstrated three significantly elevated meta-PMR values that included skin (1.69), malignant melanoma (2.25), and multiple myeloma (1.42). There was one significantly elevated metarelative risk for esoph-
ageal cancer (2.03). Incidence studies showed significant meta-SIR for cancers of the stomach (1.58), prostate (1.29), and testis (1.83).

As shown in Table 3, only one cancer type, non-Hodgkin lymphoma, had two mortality OR analyses, and both were significant. The estimated mOR was essentially based on Ma et al ${ }^{14}$ due to the much larger sample size of firefighters ( $n=4800$ ) compared with 23 for Figgs et al. ${ }^{15}$ Odds ratios were significantly higher for buccal cavity/ pharynx (5.90) and Hodgkin's disease $(2.4)^{14}$ as well as the single incidence study related to bladder cancer (2.11) and non-Hodgkin's lymphoma (3.27). ${ }^{22}$
The next step was to determine the likelihood of cancer risk based on the three criteria assessment. Cancers receiving "probable" and "possible" designations are shown in Table 4. Based on evaluating the first criterion "pattern of metarelative risk" for the 20 cancer sites, eight were designated as "probable," four as "possible," and eight as an unlikely risk. Based on the second criteria "study type" stomach, rectum, skin cancer, and malignant melanoma risk were downgraded because of reliance on PMR studies for statistical significance or the mSMR point risk estimate was not significant and $\leq 1.1$.
For the third criterion, "inconsistency" among all studies caused a downgrading for only colon cancer to "possible." This inconsistency may have been related to several factors, including study type and a cohort effect. There were 14 SMR and PMR colon cancer studies with elevated meta-risk estimates of 1.34 and 1.25 , respectively (Table 2). Of these 14 studies, there were 11 (78.6\%) with firefighters employed on or before 1950. In contrast, there were six mRR and SIR studies with meta-risk estimates of 0.91 and 0.90 , respectively, with half employed on or before 1950. It is possible that the older cohorts had higher exposures due to a lack of aware-
ness of the hazards or use of protective equipment.

A final check on the three criteria assessment presented in Table 4 was made by calculating an overall summary of cancer risk across all studies (ie, SMR, PMR, RR, SIR, OR). There was agreement that cancer was unlikely between the criteria assessment and the not significant summary risk estimates for esophagus, liver, pancreas, larynx, lung, bladder, kidney, and Hodgkin's disease and all cancers (Table 5). Differences between the two approaches were found for cancers of the buccal cavity/pharynx and leukemia because these were designated as possible by the criteria assessment but as not significant in the summary risk estimate. The remaining cancers were all rated as probable or possible and all had significant summary risk estimates. Of note, testicular cancer received the highest summary risk estimate ( $\mathrm{OR}=2.02 ; 95 \% \mathrm{CI}=$ $1.30-3.13$ ) related to the SIR studies compared with the "possible" designation by the three criteria assessment.

## Discussion

The meta-analysis and criteria assessment designate the likelihood of cancer among firefighters as probable for multiple myeloma and prostate cancer. Thus, the findings related to multiple myeloma are in agreement with Howe and Burch. ${ }^{4}$ The Philadelphia firefighter study ${ }^{13}$ was the largest cohort study reported to date investigating exposureresponse relationships. For Philadelphia firefighters, the SMR results for multiple myeloma demonstrated an increasing trend with duration of employment as a firefighter: 0.73 ( $95 \%$ $\mathrm{CI}=0.10-5.17$ ) for under 9 years, $1.50(95 \% \mathrm{CI}=0.48-4.66)$ for 10 to 19 years, and $2.31(95 \% \mathrm{CI}=1.04-$ 5.16) with six observed deaths for greater than 20 years. Except for race, there are essentially no known risk factors for multiple myeloma other than occupational exposures (eg, paints, herbicides, insecticides,
TABLE 1
Characteristics of Studies From Electronic Search

| Reference | Company Location | Design/Analysis | Study Period | Number of Workers | Comparison Group | Exposure Variable | Exposure Source | Cancer Source | Cofactors |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Baris, $2001{ }^{13}$ | Philadelphia | Cohort mortality (SMR) | 1925-1986 | 7789 | INT/NGP/NED | 1, 3, 5 | ER | DC | Age |
| Ma, $1998{ }^{14}$ | 24 US states | Case-control (MOR) | 1984-1993 | 6607 | INT | 4 | DC | DC | Age/race |
| Figgs, $1995{ }^{15}$ | 24 US states | Case-control (MOR) | 1984-1989 | 23890 (cases) | RGP | 4 | DC | DC | Age |
| 119,450 (controls) |  |  |  |  |  |  |  |  |  |
| Burnett, 1994 ${ }^{16}$ | 27 US states | PMR | 1984-1990 | 5744 | INT | 4 | DC | DC | Age |
| Demers, $1993{ }^{17}$ | 4 US states | Case-control (OR) | 1977-1981 | 692 (cases) | LGP | 4 | TRV | TRV | Age |
| 1683 (controls) |  |  |  |  |  |  |  |  |  |
| Demers, 1992a ${ }^{18}$ | Seattle, Tacoma (WA) | Cohort mortality (SMR) | 1944-1979 | 4528 |  | 4 | ER | DCN, TRV | Age |
|  |  | Incidence (SIR) |  |  | INT/LW/NGP |  |  |  |  |
| Demers, 1992b ${ }^{19}$ | Seattle, Tacoma, WA Portland | Cohort mortality (SMR) | 1944-1979 | 4546 | INT/LW/NGP | 2, 3 | ER | DCN | Age |
| Beaumont, $1991{ }^{20}$ | San Francisco | Cohort mortality (RR) | 1940-1970 | 3066 | NGP | 3, 6 | ER | DCN | Age/yr |
| Grimes, 1991 ${ }^{21}$ | Honolulu | PMR, RR | 1969-1988 | 205 | RGP | 3, 4 | ER | DC | Race |
| Sama, 1990 ${ }^{22}$ | Massachusetts | Case-control (MOR) | 1982-1986 | 315 | LW/RGP | 4,7 | TRV | TR | Age/smoke |
| Vena, $1987^{23}$ | Buffalo | Cohort mortality (SMR) | 1950-1979 | 1867 | NGP | 3 | ER | DCN | Age/yr |
| Feuer, 1986 ${ }^{24}$ | New Jersey | PMR | 1974-1980 | 263 | LW/RGP/NGP | 3, 8 | ER | DCN | Age |
| Morton, 1984 ${ }^{25}$ | Portland, Vancouver | Incidence (SIR) | 1963-1977 | 1678 | RGP | 4 | TR | TRV | Age |
| Dubrow, 1983 ${ }^{26}$ | British \& USA | Cohort mortality (SMR) | 1950-1977 | - | - | 4 | AR | DC | None |
| Musk, 1978 ${ }^{27}$ | US | Cohort mortality (SMR) | 1915-1975 | 5655 | RGP, NGP | 4 | ER | DC | Age |
| Berg $1975{ }^{28}$ | US, Great Britain | Cohort mortality (SMR) | $\begin{aligned} & \text { 1949-1953 } \\ & \text { and } \end{aligned}$ | - | NGP | 4 | DC | DC | Age |
| 促 |  |  |  |  |  |  |  |  |  |
| Stang, 2003 ${ }^{29}$ | Germany | Case-control OR) | 1995-1997 | 269 (cases) | RGP | 4 | ER | MR | Age |
| 797 (controls) |  |  |  |  |  |  |  |  |  |
| Bates, 2001 ${ }^{30}$ | New Zealand | Cohort mortality (SMR) Incidence (SIR) | 1977-1995 | 4221 | NGP | 3 | AR | DC, TR | Age/yr |
| Firth, 1996 ${ }^{31}$ | New Zealand | Incidence (SIR) | 1972-1984 | 26207 | NED | 4 | TR | TR | Age |
| Deschamps $1995{ }^{32}$ | France | Cohort mortality (SMR) | 1977-1991 | 830 | NGP | 2 | ER | DCN | Age |
| Delahunt, 1995 ${ }^{33}$ | New Zealand | Case-control (RR) | 1978-1986 | 710 (cases) | NGP | 4 | TR | TR | Age/smoke |
| 12,756 (controls) |  |  |  |  |  |  |  |  |  |
| Aronson, 1994 ${ }^{34}$ | Canada | Cohort mortality (SMR) | 1950-1989 | 5414 | RGP | 3, 6, 7 | ER | DCN | Age/yr |
| Tornling, 1994 ${ }^{35}$ | Sweden | Cohort mortality (SMR) | 1931-1983 | 1153 | LGP | 1, 3, 7 | ER | DC, TR | Age/yr |
|  |  | Incidence (SIR) |  |  |  |  |  |  |  |
| Giles, $1993{ }^{36}$ | Australia | Incidence (SIR) | 1980-1989 | 2865 | RGP | 3, 6, 7 | TRV | TR | Age |
| Guidotti, $1993{ }^{37}$ | Canada | Cohort mortality (SMR) | 1927-1987 | 3328 | RGP | 2 | ER | DCN | Age/yr |
| Hansen, 1990 ${ }^{38}$ | Denmark | Cohort mortality (SMR) | 1970-1980 | 886 | NED | 4 | OTH | DC | Age (Continued) |

TABLE 1
Continued

engine exhausts, and organic solvents). ${ }^{52-57}$ Benjamin et $\mathrm{al}^{58}$ reported that blacks compared with whites have at least double the risk of being diagnosed with multiple myeloma and twice the mortality rate. Race may be ruled out as a potential factor among firefighters, because cancer risk was investigated primarily for whites.

The analyses for non-Hodgkin's lymphoma were consistent across a diversity of study designs, including SMR, PMR, SIR, and OR incident/ mortality studies. All showed elevated meta-risk or point estimates. The overall summary risk estimate was significantly elevated at 1.51 ( $95 \% \mathrm{CI}=1.31-1.73$ ). Hence, nonHodgkin's lymphoma is considered a probable cancer risk for firefighters. Non-Hodgkin's lymphoma is, however, several cancer types with five International Classification of Disease (ICD) codes (200, 202.0, 202.1, 202.8, 202.9). Of importance is how the definition of non-Hodgkin's lymphoma by ICD code may contribute to the variability in study findings. For example, in a study by Demers et $\mathrm{al}^{19}$ comparing firefighters with police, the mortality incidence density ratio for "lymphosarcoma and reticulosarcoma" (ICD 200) was not elevated $(0.81)^{19}$ but was (1.40) for "other lymphatic/hematopoietic" (ICD 202, 203). Subsequent to the time period covered in this review, Ma et al ${ }^{59}$ examined Florida firefighters but evaluated only one of two cancers for ICD code 200, ie, lymphosarcoma but not reticular sarcoma and found nonsignificance $(S M R=0.94)$. Hence, these studies demonstrate the importance of being cognizant that differences in cancer risk estimates and interpretation of risk may be influenced by outcome definition.

Results showing a probable association for prostate cancer is curious. Prostate cancer is the most common malignancy affecting men and is the second leading cause of cancer. ${ }^{60}$ Risk of developing prostate cancer is associated with advancing age, black

TABLE 2
Metarelative Risk Estimates and Test for Inconsistency for Mortality and Incidence*

| Disease | Number of Studies | Reference | Observed | Expected | Metarelative Risk | 95\% <br> Confidence Interval | $P$ Value Inconsistency |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Mortality studies |  |  |  |  |  |  |  |
| Standardized mortality ratio (SMR) |  |  |  |  |  |  |  |
| All causes (001-999) | 12 | $\begin{aligned} & 13,19,23,27,30, \\ & 32,34 \\ & 35,37-40 \end{aligned}$ | 8384 | 9273.8 | 0.90 | 0.85-0.97 | $<0.00$ |
| All cancers (140-209) | 13 | $\begin{aligned} & 13,19,23,27,30, \\ & 32,34 \\ & 35,37-40,51 \end{aligned}$ | 1801 | 1799.9 | 1.00 | 0.93-1.08 | 0.02 |
| Buccal cavity and pharynx (140-149) | 5 | 13, 19, 32, 34, 37 | 34 | 29.8 | 1.14 | 0.79-1.60 | 0.84 |
| Esophagus (150) | 4 | 13, 19, 23, 34 | 17 | 25.1 | 0.68 | 0.39-1.08 | 0.62 |
| Stomach (151) | 7 | $\begin{aligned} & 13,19,23,30,34, \\ & 35,37 \end{aligned}$ | 75 | 81.3 | 0.92 | 0.73-1.16 | 0.72 |
| Colon (153) | 10 | $\begin{aligned} & 13,19,23,26,28 \\ & 30,34,35,37,51 \end{aligned}$ | 252 | 188.3 | 1.34 | 1.01-1.79 | $<0.00$ |
| Rectum (154) | 6 | 13, 19, 23, 30, 34, 35 | 54 | 40.7 | 1.33 | 1.00-1.73 | 0.43 |
| Liver/gallbladder (155-156) | 5 | 13, 19, 23, 34, 35 | 22 | 21.9 | 1.00 | 0.63-1.52 | 0.92 |
| Pancreas (157) | 6 | 13, 19, 23, 34, 35, 37 | 63 | 64.2 | 0.98 | 0.75-1.26 | 0.58 |
| Larynx (161) | 3 | 13, 19, 34 | 8 | 13.7 | 0.58 | 0.25-1.15 | 0.82 |
| Lung (162) | 8 | $\begin{aligned} & 13,19,30,34,35,37 \\ & 38,51 \end{aligned}$ | 378 | 359.2 | 1.05 | 0.95-1.16 | 0.50 |
| Skin (173) | 3 | 13, 19, 37 | 16 | 15.7 | 1.02 | 0.58-1.66 | 0.68 |
| Malignant melanoma (172) | 2 | 30, 34 | 4 | 5.9 | 0.67 | 0.18-1.70 | 0.23 |
| Prostate (185) | 6 | 13, 19, 23, 34, 35, 37 | 104 | 91 | 1.14 | 0.93-1.39 | 0.67 |
| Testis (186) | 1 | 34 | 3 | 1.2 | 2.50 | 0.50-7.30 | - |
| Bladder (188) | 6 | 13, 19, 23, 30, 34, 37 | 41 | 33.0 | 1.24 | 0.68-2.26 | 0.03 |
| Kidney (189) | 6 | 13, 19, 23, 34, 35, 37 | 30 | 30.9 | 0.97 | 0.44-2.13 | 0.01 |
| Brain and nervous system (191-192) | 8 | $\begin{aligned} & 13,19,23,27,30,34 \\ & 35,37 \end{aligned}$ | 64 | 46.1 | 1.39 | 0.94-2.06 | 0.07 |
| Non-Hodgkin's lymphoma $(200,202)$ | 3 | 13, 19, 34 | 30 | 20.6 | 1.46 | 0.98-2.08 | 0.92 |
| Hodgkin's disease (201) | 2 | 19, 34 | 4 | 5.1 | 0.78 | 0.21-2.01 | 0.59 |
| Multiple myeloma (203) | 4 | 13, 26, 34, 51 | 24 | 14.2 | 1.69 | 1.08-2.51 | 0.15 |
| Leukemia (204-208) | 2 | 13, 19 | 30 | 29.9 | 1.00 | 0.68-1.43 | 0.27 |
| Proportional mortality ratio (PMR) |  |  |  |  |  |  |  |
| All cancers (140-209) | 6 | 16, 24, 39, 48, 49, 50 | 2443 | 2215.7 | 1.10 | 1.06-1.15 | 0.64 |
| Buccal cavity and pharynx (140-149) | - |  | - | - | - | - | - |
| Esophagus (150) | - |  | - | - | - | - | - |
| Stomach (151) | - |  | - | - | - | - | - |
| Colon (153) | 4 | 28, 48, 49, 50 | 99 | 79.2 | 1.25 | 0.90-1.74 | 0.08 |
| Rectum (154) | 1 | 16 | 37 | 25 | 1.48 | 1.05-2.05 | - |
| Liver/gallbladder (155-156) | - |  | - | - | - | - | - |
| Pancreas (157) | - |  | - | - | - | - | - |
| Larynx (161) | - |  | - | - | - | - | - |
| Lung (162) | 4 | 16, 48, 49, 50 | 773 | 742.1 | 1.04 | 0.88-1.23 | 0.04 |
| Skin (172-173) | 2 | 16, 24 | 42 | 24.8 | 1.69 | 1.22-2.29 | 0.41 |
| Malignant melanoma (172) | 2 | 48, 49 | 9 | 4 | 2.25 | 1.03-4.27 | 0.49 |
| Prostate (185) | - |  | - | - | - | - | (Continued) |

TABLE 2
Continued

| Disease | Number of Studies | Reference | Observed | Expected | Metarelative Risk | ```Confidence Interval``` | $P$ Value Inconsistency |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Testis (186) | - |  | - | - | - | - | - |
| Bladder (188) | 1 | 16 | 37 | 37.4 | 0.99 | 0.70-1.37 | - |
| Kidney (189) | 1 | 16 | 53 | 36.8 | 1.44 | 1.08-1.89 | - |
| Brain and nervous system (191-192) | 4 | 16, 48, 49, 50 | 64 | 54.9 | 1.17 | 0.90-1.49 | 0.27 |
| Non-Hodgkin's lymphoma $(200,202)$ | 1 | 16 | 66 | 50 | 1.32 | 1.02-1.67 | - |
| Hodgkin's disease (201) | - |  | - | - | - | - | - |
| Multiple myeloma (203) | 4 | 16, 48, 49, 50 | 46 | 32.5 | 1.42 | 1.04-1.89 | 0.88 |
| Leukemia (204-208) | 2 | 16, 24 | 65 | 53.5 | 1.21 | 0.94-1.55 | 0.47 |
| Relative risk (RR) |  |  |  |  |  |  |  |
| All causes (001-999) | - | - | - | - | - | - | - |
| All cancers (140-209) | 2 | 20, 21 | 291 | 295.6 | 0.98 | 0.87-1.10 | 0.17 |
| Buccal cavity and Pharynx (140-149) | 1 | 20 | 11 | 7.7 | 1.43 | 0.71-2.57 | - |
| Esophagus (150) | 1 | 20 | 12 | 5.9 | 2.03 | 1.05-3.57 | - |
| Stomach (151) | 2 | 20, 21 | 25 | 20.6 | 1.21 | 0.80-1.81 | 0.55 |
| Colon (153) | 2 | 20, 21 | 25 | 27.5 | 0.91 | 0.60-1.36 | 0.92 |
| Rectum (154) | 1 | 20 | 13 | 9 | 1.44 | 0.77-2.49 | - |
| Liver (155-156) | - | - | - | - | - | - | - |
| Pancreas (157) | 1 | 20 | 17 | 13.6 | 1.25 | 0.73-2.00 | - |
| Larynx (161) | 1 | 20 | 3 | 3.8 | 0.79 | 0.17-2.35 | - |
| Lung (162) | 1 | 20 | 60 | 71.4 | 0.84 | 0.64-1.08 | - |
| Skin (172-173) | 1 | 20 | 7 | 4.1 | 1.71 | 0.68-3.49 | - |
| Malignant melanoma (172) | - | - | - | - | - | - | - |
| Prostate (185) | 2 | 20, 21 | 19 | 24.3 | 0.78 | 0.13-4.82 | $<0.00$ |
| Testis (186) | - | - | - | - | - | - | - |
| Bladder (188) | - | - - | - | - | - | - | - |
| Kidney (189) | 1 | 20 | 4 | 5.9 | 0.68 | 0.19-1.74 | - |
| Brain and nervous system (191-192) | 2 | 20, 21 | 9 | 7.1 | 1.26 | 0.55-2.34 | 0.14 |
| Non-Hodgkin's lymphoma $(200,202)$ | - | - | - | - | - | - | - |
| Hodgkin's disease (201) | - | - | - | - | - | - | - |
| Multiple myeloma (203) | - | - - | - | - | - | - | - |
| Leukemia (204-208) | 1 | 20 | 6 | 9.8 | 0.61 | 0.22-1.33 | - |
| Incidence studies (SIR) |  |  |  |  |  |  |  |
| All cancers (140-209) | 3 | 30, 35, 36 | 367 | 366.6 | 1.00 | 0.90-1.11 | 0.61 |
| Buccal cavity and pharynx (140-149) | 2 | 18, 36 | 25 | 19.6 | 1.28 | 0.83-1.88 | 0.73 |
| Esophagus (150) | 2 | 18, 30 | 10 | 7.6 | 1.32 | 0.63-2.42 | 0.51 |
| Stomach (151) | 3 | 18, 30, 35 | 38 | 24.1 | 1.58 | 1.12-2.16 | 0.33 |
| Colon (153) | 4 | 18, 30, 35, 36 $\dagger$ | 59 | 65.3 | 0.9 | 0.69-1.17 | 0.37 |
| Rectum (154) | 3 | 18, 30, 35 | 41 | 36.1 | 1.14 | 0.81-1.54 | 0.4 |
| Liver (155-156) | 1 | 35 | 4 | 4.7 | 0.85 | 0.23-2.18 | - |
| Pancreas (157) | 4 | 18, 30, 35, 36 | 22 | 18.2 | 1.21 | 0.76-1.83 | 0.83 |
| Larynx (161) | 2 | 18, 31 | 13 | 8.3 | 1.57 | 0.17-14.51 | $<0.00$ |
| Lung (162) | 4 | 18, 30, 35, 36 | 111 | 120.0 | 0.93 | 0.76-1.11 | 0.83 |
| Skin (172-173) | 1 | 35 | 5 | 3.3 | 1.52 | 0.49-3.54 | - |
| Malignant melanoma (172) | 4 | 18, 30, 35, 36 | 60 | 47.9 | 1.25 | 0.96-1.61 | 0.87 |
| Prostate (185) | 4 | 18, 30, 35, 36 | 147 | 114.1 | 1.29 | 1.09-1.51 | $\begin{gathered} 0.56 \\ \text { (Continued) } \end{gathered}$ |

TABLE 2
Continued

| Disease | $\begin{array}{c}\text { Number of } \\ \text { Studies }\end{array}$ | Reference | Observed | Expected | $\begin{array}{c}\text { Metarelative } \\ \text { Risk }\end{array}$ | $\begin{array}{c}95 \% \\ \text { Confidence } \\ \text { Interval }\end{array}$ |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- | \(\left.\begin{array}{c}\boldsymbol{P} Value <br>

Inconsistency\end{array}\right]\)

Note. Codes of the International Classification of Causes of Death (9th Revision) in parentheses; published data for references 48-50 in Howe and Birch. ${ }^{4}$
*Meta analysis completed only for two or more studies.
$\dagger$ Reference 36 is a combination of colon and rectum cancers.
table 3
Mortality and Incidence Studies for Case-Control/Mortality Odds Ratio Studies


[^1]| TABLE 4 |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Likelihood of Cancer Risk Among Firefighters After Employing Pattern of Metarelative Risk Association, Study Type, and Inconsistency Among Studies |  |  |  |  |  |  |  |  |  |  |  |
|  |  |  |  |  |  |  |  |  |  |  |  |
| Cancer Site | Pattern of Metarelative Risk Association |  |  |  |  |  | Likelihood of Cancer Risk | Criteria 2 |  | Criteria 3 |  |
|  | mSMR | mPMR | mSMR and PMR | mRR | mSIR | mOR |  | Study Type | Likelihood of Cancer Risk | Inconsistency | Likelihood of Cancer Risk |
| Buccal | + | NA | NC | NC | + | - | Possible | No change | Possible | No change | Possible |
| Stomach | - | NA | NC | + | + + | - | Probable | Down one | Possible | No change | Possible |
| Colon | + + | $+$ | + + | - | - | - | Probable | No change | Probable | Down one | Possible |
| Rectum | + | NC | + + | NC | + | - | Probable | Down one | Possible | No change | Possible |
| Skin | - | ++ | ++ | NC | NC | - | Probable | Down one | Possible | No change | Possible |
| Malignant melanoma | - | + + | $-$ | NA | + | - | Probable | Down one | Possible | No change | Possible |
| Prostate | + | NA | NC | - | + + | - | Probable | No change | Probable | No change | Probable |
| Testis | NC | NA | NC | NA | + + | - | Possible | No change | Possible | No change | Possible |
| Brain | + | $+$ | $+$ | $+$ | $+$ | - | Possible | No change | Possible | No change | Possible |
| Non-Hodgkin's lymphoma | + | NC | ++ | NA | NC | + + | Probable | No change | Probable | No change | Probable |
| Multiple myeloma | + + | + + | ++ | NA | NA | - | Probable | No change | Probable | No change | Probable |
| Leukemia | - | $+$ | + | NC | $+$ | - | Possible | No change | Possible | No change | Possible |

 is $\leq 1.1$ and not significant at the $5 \%$ level.
NA indicates no available studies; NC, not able to calculate because only one study of that type available.
Study type: down one level, the metarelative risk $(++)$ is based primarily on mPMR studies and/or negative ( - ) mSMR studies.
Inconsistency among studies: down one level heterogeneity significant among all combined studies at the $10 \%$ level.
ethnicity, a positive family history, and may be influenced by diet. Although the positive association with prostate cancer may be due to some of these factors, it is unlikely that these entirely explain the findings; most studies analyzed white men adjusting for age. The summary risk estimate was 1.28 ( $95 \% \mathrm{CI}=1.15-$ 1.43). The mSIR was significantly elevated, and all individual studies showed excess SIR values. Parent and Siemiatycki, ${ }^{61}$ in a review article, concluded that there was suggestive epidemiologic evidence for prostate cancer associated with exposure to pesticides and herbicides, metallic dusts, metal working fluids, polycyclic aromatic hydrocarbon, and diesel engine emissions. Certainly firefighters are exposed to these latter two agents. Recently, exposure to complex mixture in the semiconductor industry also has been associated with an increase in prostate cancer. ${ }^{62}$ Thus, it is possible that some of the mixed exposures experienced by firefighters may be prostate carcinogens. Ross and Schottenfeld ${ }^{63}$ have cautioned, however, against associating occupational exposures with prostate cancer.

Although there were only four studies evaluating testicular cancer, we propose upgrading the likelihood of cancer risk from possible to probable. This upgrade is suggested because testicular cancer had the largest summary point estimate ( $2.02,95 \% \mathrm{CI}=$ $1.30-3.13$ ) as well as consistency among the one SMR study, two incidence studies, and one casecontrol study showing elevated risk estimates between 1.15 and 4.30. Testicular cancer is the most common malignancy between the ages of 20 and 34. Except for cryptorchism, no risk factor has been clearly demonstrated. ${ }^{64}$ Because testicular cancer occurs among younger men with high survival, mortality studies are less germane. Bates et al ${ }^{30}$ showed an increase in the incident cases of testicular cancer with firefighter exposure duration as follows: 10 years:

TABLE 5
Summary of Likelihood of Cancer Risk and Summary Risk Estimate ( $95 \% \mathrm{Cl}$ ) Across All Types of Studies for All Cancers

| Cancer Site | Likelihood of Cancer Risk by Criteria | Summary Risk Estimate (95\% CI) | Comments |
| :---: | :---: | :---: | :---: |
| Multiple myeloma | Probable | 1.53 (1.21-1.94) | Consistent with mSMR and PMR (1.50, $95 \% \mathrm{CI}=1.17-1.89)$ Based on 10 analyses <br> Heterogeneity—not significant at the $10 \%$ level |
| Non-Hodgkin lymphoma | Probable | 1.51 (1.31-1.73) | Only two SMR and another PMR studies <br> Slightly higher than mSMR and PMR (1.36, 95\% CI = 1.10-1.67) <br> Based on eight analyses <br> Heterogeneity-not significant at the 10\% level |
| Prostate | Probable | 1.28 (1.15-1.43) | Consistent with $\operatorname{mSIR}(1.29,95 \% \mathrm{CI}=1.09-1.51)$ <br> Based on 13 analyses <br> Heterogeneity—not significant at the 10\% level |
| Testis | Possible | 2.02 (1.30-3.13) | Slightly higher than $\mathrm{mSIR}(1.83,95 \% \mathrm{CI}=1.13-2.79)$ <br> Based on four analyses <br> Heterogeneity-not significant at the 10\% level |
| Skin | Possible | 1.39 (1.10-1.73) | Slightly lower than mSMR and PMR $(1.44,95 \% \mathrm{CI}=1.10-1.87)-$ derived on basis of PMR studies <br> Based on eight analyses <br> Heterogeneity—not significant at the 10\% level |
| Malignant melanoma | Possible | 1.32 (1.10-1.57) | Slightly higher than mSMR and PMR (1.29, 95\% CI $=0.68-2.20)$ Based on 10 analyses <br> Heterogeneity—not significant at the 10\% level |
| Brain | Possible | 1.32 (1.12-1.54) | Slightly higher than mSMR and PMR (1.27, 95\% CI $=0.98-1.63$ ) Based on 19 analyses <br> Heterogeneity—not significant at the 10\% level; there was heterogeneity among SMR studies |
| Rectum | Possible | 1.29 (1.10-1.51) | Slightly lower than mSMR and PMR (1.39, 95\% CI = 1.12-1.70) Based on 13 analyses <br> Heterogeneity—not significant at the 10\% level |
| Buccal cavity and pharynx | Possible | 1.23 (0.96-1.55) | Slightly higher than mSMR (1.18, $95 \% \mathrm{CI}=0.81-1.66$ ) Based on nine analyses <br> Heterogeneity—not significant at the 10\% level |
| Stomach | Possible | 1.22 (1.04-1.44) | Lower than $\operatorname{mSIR}(1.58,95 \% \mathrm{Cl}=1.12-2.16)$; <br> Based on 13 analyses <br> Heterogeneity—not significant at the 10\% level |
| Colon | Possible | 1.21 (1.03-1.41) | Slightly lower than mSMR and PMR (1.31, $95 \% \mathrm{CI}=1.08-1.59$ ) Based on 25 analyses <br> Heterogeneity—significant at the 10\% level; there were heterogeneity among SMR and PMR studies |
| Leukemia | Possible | 1.14 (0.98-1.31) | Similar to mSMR and PMR (1.14, 95\% CI = 0.92-1.39) <br> Based on eight analyses <br> Heterogeneity—not significant at the $10 \%$ level |
| Larynx | Unlikely | 1.22 (0.87-1.70) | Higher than mSMR $(0.58,95 \% \mathrm{CI}=0.25-1.15)$ <br> Based on seven analyses <br> Heterogeneity-not significant at the 10\% level |
| Bladder | Unlikely | 1.20 (0.97-1.48) | Similar to mSMR and $\operatorname{PMR}(1.24,95 \% \mathrm{CI}=0.83,1.49)$ Based on 11 analyses <br> Heterogeneity-significant at the 10\% level; there was heterogeneity among SMR studies |
| Esophagus | Unlikely | 1.16 (0.86-1.57) | Higher than mSMR ( $0.68,95 \% \mathrm{CI}=0.39-1.08$ ) <br> Based on eight analyses <br> Heterogeneity—not significant at the 10\% level |
| Pancreas | Unlikely | 1.10 (0.91-1.34) | Slightly higher than mSMR ( $0.98,95 \% \mathrm{CI}=0.75-1.26$ ) <br> Based on 13 analyses <br> Heterogeneity—not significant at the 10\% level |
| Kidney | Unlikely | 1.07 (0.78-1.46) | Similar to mSMR and PMR (1.23, 95\% CI = 0.94-1.59) Based on 12 analyses <br> Heterogeneity-significant at the 10\% level; there was heterogeneity among SMR studies |

TABLE 5
Continued

| Cancer Site | Likelihood of Cancer <br> Risk by Criteria | Summary Risk <br> Estimate (95\% CI) | Comments |
| :--- | :---: | :---: | :--- |

Cl indicates confidence interval; SMR, standardized mortality ratio; PMR, proportional mortality ratio; SIR, standardized incidence ratio.
$\operatorname{SIR}=1.39,95 \% \mathrm{CI}=0.2-5.0 ; 11$ to 20 years: SIR $=4.03,95 \% \mathrm{CI}=$ 1.3-9.4. In those exposed greater than 20 years, the risk estimate remained elevated but declined (SIR $=$ $2.65,95 \% \mathrm{CI}=0.3-9.6$ ), possibly because testicular cancer generally occurs at a younger age. Bates et al ${ }^{30}$ argued that, although the reason for the excess risk of testicular cancer remained obscure, the possibility that this is a chance finding was low because incident studies are likely the most appropriate methodology for a cancer that can be successfully treated.

The 1990 findings of Howe and Burch ${ }^{4}$ showing a positive association with brain cancer and malignant melanoma are compatible with our results because both had significant summary risk estimates. Brain cancers were initially scored as probable but then downgraded to possible (Table 5). There was inconsistency among the SMR studies, which resulted in the use of the randomeffects model, yielding confidence limits that were not significant $(\mathrm{SMR}=1.39,95 \% \mathrm{CI}=0.94-2.06)$ (Table 2). This inconsistency primarily resulted from the Baris et al study, ${ }^{13}$ a 61 -year follow up of 7789 firefighters demonstrating a marked reduction in brain cancer (SMR $=$ $0.61,95 \% \mathrm{CI}=0.31-1.22)$. As
noted in Table 4, however, there were elevated, but not significant, risk estimates across all studies, ie, mSMR, mPMR, mRR, and mSIR. This consistency is all the more remarkable given the diversity of rare cancers included in the category "brain and nervous system." Furthermore, there was a 2003 study by Krishnan et al ${ }^{65}$ published after our search that examined adult gliomas in the San Francisco Bay area of men in 35 occupational groups. This study showed that male firefighters (six cases and one control) had the highest risk with an odds ratio of 5.93 , although the confidence intervals were wide and not significant. In addition, malignant melanoma was also initially scored as probable but was downgraded to "possible" due to study type. This study downgrade was related to the negative SMR (-) and reliance primarily on a PMR study. Thus, in conclusion, our study supports a probable risk for multiple myeloma, similar to Howe and Burch's ${ }^{4}$ findings, and a possible association with malignant melanoma and brain cancer.

## Summary

We implemented a qualitative three-criteria assessment in addition to the quantitative meta-analyses. Based on the more traditional quan-
titative summary risk estimates shown in Table 5, 10 cancers, or half, were significantly associated with firefighting after the three cancers were designated as a probable risk based on the quantitative meta-risk estimates and our three criteria assessment. These cancers included multiple myeloma, non-Hodgkin's lymphoma, and prostate. A recommendation is also made, however, for upgrading testicular cancer to "probable" based on the twofold excess summary risk estimate and the consistency among the studies. Thus, firefighter risk for these four cancers may be related to the direct effect associated with exposures to complex mixtures, the routes of delivery to target organs, and the indirect effects associated with modulation of biochemical or physiologic pathways. In anecdotal conversations with firefighters, they report that their skin, including the groin area, is frequently covered with "black soot." It is noteworthy that testicular cancer had the highest summary risk estimate (2.02) and skin cancer had a summary risk estimate (1.39) higher than prostate (1.28). Certainly, Edelman et al ${ }^{3}$ at the World Trade Center, although under extreme conditions, revealed the hazards that firefighters may encounter only because air monitoring was performed.

As noted in Table 1, approximately half of the studies used local, regional, or national general population rates as the comparison group. These general population comparison groups raise concern that the actual risk of cancer may be underestimated due to the healthy worker effect related to the strict physical entry requirements, maintenance of better physical fitness, and good health benefits. The healthy worker bias may be less pronounced, however, for cancer than for conditions such as coronary heart disease. Furthermore, tobacco is unlikely a contributing factor because cancers known to be associated with smoking such as lung, bladder, and larynx were designated as unlikely and corresponding summary risk estimates were not statistically significant.

These findings of an association of firefighting with significant increased risk for specific types of cancer raise red flags and should encourage further development of innovative comfortable protective equipment allowing firefighters to do their jobs without compromising their health. Studies are especially needed that better characterize the type and extent of exposures to firefighters.

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[^1]:    *Two control groups available; police rather than state employees selected as most comparable. Significance difference only for malignant melanoma when using state employees odds ratio and 95\% confidence interval was 2.92 (1.70-5.03).
    $\dagger$ Mortality odds ratio (mOR) calculated only for non-Hodgkin lymphoma as only case-control study with at least two studies. mOR estimated based primarily on larger sample in Ma et al. ${ }^{14}$

